

PROPHET INEQUALITY:

Def: A stopping rule τ w.r.t. random variables X_1, \dots, X_n is a r.v. τ w/ values in $\{1, 2, \dots, n\}$ and the property that for all $t \in \{1, \dots, n\}$ the occurrence or non-occurrence of event $\tau = t$ only depends on the values of X_1, \dots, X_t .

[Krengel-Sucheston - Garling '79]: If X_1, \dots, X_n are independent, non-negative there exists a stopping rule τ w.r.t. X_1, \dots, X_n such that

$$\mathbb{E}[X_\tau] \geq \frac{1}{2} \mathbb{E}[\max_i X_i]$$

Proof: Let F_1, \dots, F_n be the distn's of X_1, \dots, X_n

- Let $\tau(\beta) = \arg \min_t \{X_t \geq \beta\}$ for some β TBD

(i.e. stop the first time you see a sample $\geq \beta$)

- Let $q_\beta(\beta) = \Pr[X_t < \beta, \forall t] = \prod_t F_t(\beta^-)$

- $\mathbb{E}[X_{\tau(\beta)}] \geq$

$$\beta \cdot \Pr[\exists t, t' \text{ s.t. } X_t \geq \beta, X_{t'} \geq \beta] + \sum_{t=1}^n \mathbb{E}[X_t | X_t \geq \beta \wedge X_{t'} < \beta, \forall t' \neq t] \cdot \Pr[X_t \geq \beta \wedge X_{t'} < \beta, \forall t' \neq t]$$

(4)

$$(*) = \sum_{t=1}^n \mathbb{E}[X_t | X_t \geq j] \cdot \Pr[X_t \geq j] \cdot \prod_{t' \neq t} \Pr[X_{t'} < j]$$

$$= \sum_{t=1}^n \left(\mathbb{E}[X_t - j | X_t \geq j] + j \right) \cdot \Pr[X_t \geq j] \cdot \prod_{t' \neq t} F_{t'}(j^-)$$

$$= \sum_{t=1}^n \left(\mathbb{E}[(X_t - j)_+] \cdot \prod_{t' \neq t} F_{t'}(j^-) + j \cdot (1 - F_t(j)) \cdot \prod_{t' \neq t} F_{t'}(j^-) \right)$$

$$\geq q(j) \cdot \sum_{t=1}^n \mathbb{E}[(X_t - j)_+] + j \cdot \Pr[\exists! t \text{ s.t. } X_t \geq j]$$

• Hence: $\mathbb{E}[X_{I(j)}] \geq (1 - q(j)) \cdot j + q(j) \cdot \sum_{t=1}^n \mathbb{E}[(X_t - j)_+]$ (***)

• On the other hand,

$$\begin{aligned} \mathbb{E}[\max X_t] &= \mathbb{E}[j + \max(X_t - j)] \\ &\leq j + \mathbb{E}[\max(X_t - j)_+] \\ &\leq j + \sum_{t=1}^n \mathbb{E}[(X_t - j)_+] \end{aligned}$$
(***)

Suppose exists j s.t. $q(j) = 1 - q(j) = \frac{1}{2}$. Proof then follows from (**) and (***).

\Rightarrow If such \bar{J} doesn't exist find \bar{J} s.t.

$$q^+(\bar{J}) = \Pr[X_t < \bar{J}, \forall t] \leq \frac{1}{2} \leq \Pr[X_t \leq \bar{J}, \forall t] = q^-(\bar{J}) \quad (\text{by } *)$$

Compare the revenue from stopping rules:

$$\begin{aligned} I^+(\bar{J}) &= \arg \min_t \{X_t > \bar{J}\} \\ \text{and } I^-(\bar{J}) &= \arg \min_t \{X_t \geq \bar{J}\} \end{aligned} \quad \left. \begin{array}{l} \text{want to show} \\ \text{that at least} \\ \text{one of these} \\ \text{two guarantees} \\ \frac{1}{2} - \text{approximation} \end{array} \right\}$$

Using the following derivation as for $\mathbb{E}[X_{I^+(\bar{J})}]$

$$\mathbb{E}[X_{I^+(\bar{J})}] \geq \bar{J} \cdot \underbrace{\Pr[\exists t, t' \text{ s.t. } X_t > \bar{J}, X_{t'} > \bar{J}]}_{(*)} + \underbrace{\sum_{t=1}^n \mathbb{E}[X_t | X_t > \bar{J} \wedge X_{t'} \leq \bar{J}, \forall t' \neq t]}_{(A)} \cdot \Pr[X_t > \bar{J} \wedge X_{t'} \leq \bar{J}, \forall t' \neq t]$$

$$(*) = \sum_{t=1}^n \mathbb{E}[X_t | X_t > \bar{J}] \cdot \Pr[X_t > \bar{J}] \cdot \prod_{t' \neq t} \Pr[X_{t'} \leq \bar{J}]$$

$$= \sum_{t=1}^n \left(\mathbb{E}[X_t - \bar{J} | X_t > \bar{J}] + \bar{J} \right) \cdot \Pr[X_t > \bar{J}] \cdot \prod_{t' \neq t} F_{t'}(\bar{J})$$

$$= \sum_{t=1}^n \left(\mathbb{E}[(X_t - \bar{J})_+] \cdot \prod_{t' \neq t} F_{t'}(\bar{J}) + \bar{J} \cdot (1 - F_t(\bar{J})) \cdot \prod_{t' \neq t} F_{t'}(\bar{J}) \right)$$

$$\geq \underbrace{q^+(\bar{J})}_{\text{by } *} \cdot \sum_{t=1}^n \mathbb{E}[X_t - \bar{J}]_+ + \bar{J} \cdot \underbrace{\Pr[\exists t \text{ s.t. } X_t > \bar{J}]}_{\text{by } *}$$

- Hence: $\mathbb{E}[X_{I^+(J)}] \geq (1-q^+(J)) \cdot \boxed{\quad} + \boxed{q^+(J)} \cdot \sum_{t=1}^n \mathbb{E}[(X_t - J)_+]$ (****)

- Recall from (**) that:

$$\mathbb{E}[X_{I(J)}] \geq \boxed{(1-q(J)) \cdot \boxed{\quad}} + \boxed{q(J)} \cdot \sum_{t=1}^n \mathbb{E}[(X_t - J)_+] \quad (***)$$

and from (****) that:

$$\mathbb{E}[\max X_t] \leq \boxed{\quad} + \sum_{t=1}^n \mathbb{E}[(X_t - J)_+] \quad (****)$$

- Case analysis: - If $\boxed{\quad} \geq \sum_{t=1}^n \mathbb{E}[(X_t - J)_+]$, then it follows from (****), (**) & (****) that

$$\mathbb{E}[X_{I(J)}] \geq \frac{1}{2} \mathbb{E}[\max X_t]$$

- If $\boxed{\quad} \leq \sum_{t=1}^n \mathbb{E}[(X_t - J)_+]$, then it follows

from (****), (*****) & (****) that

$$\mathbb{E}[X_{I^+(J)}] \geq \frac{1}{2} \mathbb{E}[\max X_t]$$



Corollary to Bulow-Klemperer '96:

For any regular dist'n F and any n :

$$\mathbb{E}_{v_1, \dots, v_n \sim F} [\text{Rev(Vickrey)}] \geq (1 - \frac{1}{n}) \mathbb{E}_{v_1, \dots, v_n \sim F} [\text{Rev(Myerson)}]$$

Proof: • Suffices to show:

$$\mathbb{E}_{v_1, \dots, v_{n-1} \sim F} [\text{Rev of Myerson on } v_1, \dots, v_{n-1}] \geq (1 - \frac{1}{n}) \mathbb{E}_{v_1, \dots, v_n \sim F} [\text{Rev of Myerson on } v_1, \dots, v_n]$$

• Sample v_1, \dots, v_{n-1} and $v'_1, v'_2, \dots, v'_{n-1}, v'_n$ as follows:

- draw iid samples u_1, \dots, u_n from F
- draw random permutation π
- set $v_i = u_{\pi(i)}, \forall i=1, \dots, n-1$
- $v'_i = u_{\pi(i)}, \forall i=1, \dots, n$

• Under above sampling:

① v_1, \dots, v_{n-1} iid from F ; hence:

$$\mathbb{E}_{\vec{u}, \pi} \left[\begin{matrix} \text{Virtual Welfare of Myerson's} \\ \text{auction on } v_1, \dots, v_{n-1} \end{matrix} \right] = \mathbb{E}_{v_1, \dots, v_{n-1} \sim F} \left[\begin{matrix} \text{Revenue of Myerson's} \\ \text{auction on } v_1, \dots, v_{n-1} \end{matrix} \right]$$

② v'_1, \dots, v'_n iid from F ; hence:

$$\mathbb{E}_{\vec{u}, \tau} \left[\text{VW of Myerson's auction on } v'_1, \dots, v'_n \right] = \mathbb{E}_{v'_1, \dots, v'_n \sim F} \left[\begin{array}{l} \text{Revenue of Myerson's} \\ \text{auction on } v'_1, \dots, v'_n \end{array} \right]$$

③ For each realization of \vec{u} :

$$\mathbb{E}_{\pi} \left[\text{Virtual Welfare of Myerson's auction on } v_1, \dots, v_{n-1} \right] \geq$$

$$\left(1 - \frac{1}{n}\right) \cdot \mathbb{E}_{\pi} \left[\text{Virtual Welfare of Myerson's auction on } v'_1, \dots, v'_n \right]$$